9(·) 10.5 Function of Continuous Random Variables: SISO

Reconsider the derived random variable Y = g(X).

$$X \longrightarrow g(\cdot) \longrightarrow Y = g(X) \qquad E_X. \quad Y = X^2 \qquad g(\cdot) = (\cdot)^2$$

g(-) = 4 | - 1.5

Recall that we can find $\mathbb{E}Y$ easily by (22):

$$\mathbb{E}Y = \mathbb{E}\left[g(X)\right] = \int_{\mathbb{R}} g(x) f_X(x) dx$$

However, there are cases when we have to evaluate probability directly involving the random variable Y or find $f_Y(y)$ directly.

Recall that for discrete random variables, it is easy to find $p_Y(y)$ by adding all $p_X(x)$ over all x such that g(x) = y:

$$p_Y(y) = \sum_{x:g(x)=y} p_X(x).$$
 (23)

For continuous random variables, it turns out that we can't⁴⁵ simply integrate the pdf of X to get the pdf of Y.

10.61. For Y = g(X), if you want to find $f_Y(y)$, the following **two-step** procedure will always work and is easy to remember:

- (a) Find the cdf $F_Y(y) = P[Y \le y]$.
- (b) Compute the pdf from the cdf by "finding the derivative" $f_Y(y) = \frac{d}{dy} F_Y(y)$ (as described in 10.13).

10.62. Linear Transformation: Suppose Y = aX + b. Then, the cdf of Y is given by

$$\bigcirc_{F_Y(y)} = P\left[Y \le y\right] = P\left[aX + b \le y\right] = \begin{cases} P\left[X \le \frac{y-b}{a}\right], & a > 0, \\ P\left[X \ge \frac{y-b}{a}\right], & a < 0. \end{cases}$$

Now, by definition, we know that

$$P\left[X \le \frac{y-b}{a}\right] = F_X\left(\frac{y-b}{a}\right),$$

⁴⁵When you applied Equation (23) to continuous random variables, what you would get is 0 = 0, which is true but not interesting nor useful.

and

$$P\left[X \ge \frac{y-b}{a}\right] = P\left[X > \frac{y-b}{a}\right] + P\left[X = \frac{y-b}{a}\right]$$
$$= 1 - F_X\left(\frac{y-b}{a}\right) + P\left[X = \frac{y-b}{a}\right].$$

For continuous random variable, $P\left[X = \frac{y-b}{a}\right] = 0$. Hence,

$$F_{Y}(y) = \begin{cases} F_{X}\left(\frac{y-b}{a}\right), & a > 0, \\ 1 - F_{X}\left(\frac{y-b}{a}\right), & a < 0. \end{cases}$$

Finally, fundamental theorem of calculus and chain rule gives

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \begin{cases} \frac{1}{a} f_X\left(\frac{y-b}{a}\right), & a > 0, \\ -\frac{1}{a} f_X\left(\frac{y-b}{a}\right), & a < 0. \end{cases}$$

$$f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y-b}{a}\right), \quad a \neq 0.$$

$$Y = -5 \times +2$$

$$(24)$$

$$f_Y(y) = \left| \frac{1}{5} \right| f_X\left(\frac{y-y}{5}\right)$$

Graphically, to get the plots of f_Y , we compress f_X horizontally by a factor of a, scale it vertically by a factor of 1/|a|, and shift it to the right by b.

Of course, if a = 0, then we get the uninteresting degenerated random variable $Y \equiv b$.

Example 10.63. Suppose $X \sim \mathcal{E}(\lambda)$. Let Y = 5X Find $f_Y(y)$. $F_Y(y) = P[Y \leq y] = P[5 \times \leq y]$ $f(x) = \begin{cases} \lambda e^{-\lambda y} & e^{-\lambda y} \\ 0 & \text{otherwise} \end{cases}$ $= P[X \leq \frac{y}{5}] = F_X(\frac{y}{5})$ (hain rule)We can use (2^{4}) to get $f_Y(y)$ directly. $f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(\frac{y}{5})$ $(\frac{1}{5} \times (\frac{y}{5})) = \begin{cases} \frac{1}{5} \lambda e^{-\lambda \frac{y}{5}} & \frac{y}{5} > 0 \\ 0 & \text{otherwise} \end{cases}$ $= \begin{cases} \frac{\lambda}{5} e^{-\frac{\lambda}{5}y} & y > 0 \\ 0 & \text{otherwise} \end{cases}$ $= \begin{cases} \frac{\lambda}{5} e^{-\frac{\lambda}{5}y} & y > 0 \\ 0 & \text{otherwise} \end{cases}$ = 143

10.64. Suppose
$$(X \sim N(m, \sigma^2))$$
 but $Y = aX + b$ by some constants
a and b. Then, we can use (24) to show that $Y \sim N(am + b, \sigma^2 \sigma^2)$.

$$f_{\gamma}(y) = \frac{1}{|a|} f_{\chi}(\frac{y-b}{\sigma}) = \frac{1}{(ax^2 \sigma)|a|} e^{xp} \left(-\frac{1}{2}\sigma^2\right)^2 = \frac{y-(b+an)}{\sigma}$$

$$= \frac{1}{(ax^2 \sigma)|a|} e^{xp} \left(-\frac{1}{2}\left(\frac{y-cn}{\sigma}\right)^2\right)$$
Example 10.65. Amplitude modulation in certain communication systems can be accomplished using various nonlinear devices
such as a semiconductor diade. Suppose we model the nonlinear
device by the function $Y = X^3$ If the input X is a continuous
random variable, find the density of the output Y = X².
For $y \ge 0$, $F_{\gamma}(y) = P[X^2 \le y] = P[-F_{\gamma} \le x \le f_{\gamma}] = F_{\chi}(F_{\gamma}) - F_{\chi}(-F_{\gamma})$
 $f_{\gamma}(y) = \frac{1}{2} f_{\gamma}(x) - (-\frac{1}{4} f_{\gamma}) f_{\chi}(-F_{\gamma})$
 $f_{\gamma}(y) = \frac{1}{2} f_{\gamma}(x) - (-\frac{1}{4} f_{\gamma}) f_{\chi}(-F_{\gamma})$
 $f_{\gamma}(y) = P[X \le f_{\gamma}] = P[-f_{\gamma} \le x \le f_{\gamma}] = F_{\chi}(F_{\gamma}) - F_{\chi}(-F_{\gamma})$
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 $f_{\gamma}(y) = \frac{1}{2} f_{\gamma}(x) - (-\frac{1}{4} f_{\gamma}) f_{\chi}(-F_{\gamma})$
 $f_{\gamma}(y) = P[X \ge \frac{1}{4} f_{\gamma}] + P[-f_{\gamma}(x) - f_{\gamma}(x) + f_{\gamma}(x)] + f_{\gamma}(x) + f_{\gamma}$



Ex.
$$Y = x^2 \Rightarrow Here, g(x) = \sigma c^2$$
. () Find root(s) of $y = ze^2$:
 $\Rightarrow x = \pm /\overline{y} \quad \leftarrow \pm wo \text{ roots when } y > 0.$
(2) Find slope(s):
 $g'(x) = \frac{d}{dx} z^2 = 2\pi$
 $\frac{d}{dx} z^2 = 2\pi$
 $\frac{d}{dx}$

Ex.
$$Y = e^{X} \Rightarrow Here, g(x) = e^{A}$$
.
Assume $X \sim \mathcal{U}(-2, 2)$ (a) Suppose we want to find $\mathcal{F}_{Y}(1)$.
(a) Suppose we want to find $\mathcal{F}_{Y}(1)$.
(b) Find root: To get $Y = 1$, need
(c) $Y = 0$.
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(b) Suppose we want to find $f_{\gamma}(-1)$. () Find root: $-1 = e^{\mathbf{x}} \Rightarrow \text{No } \mathbf{x}$ satisfies Therefore, $f_{\gamma}(-1) = 0$. this.

(c) Suppose we want to find
$$f_{\gamma}(100)$$
. () Find root(s): $100 = e^{x}$
(2) Find slope(s): $g(\alpha) = e^{x} \Rightarrow g'(x) = e^{x}$.
(3) $f_{\gamma}(100) = \frac{f(x, 6)}{|e^{x, 6}|} = 0$



P[Y=1] = P[X=1] + P[X=-1]

 $= P_{X}(1) + P_{X}(-1)$ = 0.25 + 0.25 = 0.5 = $P_{Y}(1)$

 $= f_{x}(1) + f_{x}(-1)$

|slope@1| |slope@-1)



$$f_{\gamma}(2) = \frac{f_{\gamma}(-\frac{\gamma}{2})}{|g(-\frac{\gamma}{2})|} + \frac{f_{\chi}(-\frac{\gamma}{2})}{|g(-\frac{\gamma}{2})|} + \frac{f_{\chi}(-\frac{\gamma}{2})}{|g(-\frac{\gamma}{2})|} = \frac{e}{1} + \frac{e}{1} \approx 0.0507$$

$$(b) \quad f_{\gamma}(1) = \frac{f_{\chi}(-\frac{\sigma}{2})}{|g(-\frac{\gamma}{2})|} + \frac{f_{\chi}(-\frac{\gamma}{2})}{|g(-\frac{\gamma}{2})|} + \frac{f_{\chi}(-\frac{\gamma}{2})}{|g(-\frac{\gamma}{2})|} + \frac{f_{\chi}(-\frac{\gamma}{2})}{|g(-\frac{\gamma}{2})|} = \frac{e^{-\frac{\gamma}{2}}}{|g(-\frac{\gamma}{2})|} + \frac{e^{-\frac{\sigma}{2}}}{|g(-\frac{\gamma}{2})|} = \frac{e^{-\frac{\gamma}{2}}}{|g(-\frac{\gamma}{2})|} = \frac{e^{-\frac{\gamma}{2}}}{|g($$

Exercise 10.67 (F2011). Suppose X is uniformly distributed on the interval (1,2). $(X \sim \mathcal{U}(1,2).)$ Let $Y = \frac{1}{X^2}$.

- (a) Find $f_Y(y)$.
- (b) Find $\mathbb{E}Y$.

Exercise 10.68 (F2011). Consider the function

$$g(x) = \begin{cases} x, & x \ge 0\\ -x, & x < 0. \end{cases}$$

Suppose Y = g(X), where $X \sim \mathcal{U}(-2, 2)$.

Remark: The function g operates like a **full-wave rectifier** in that if a positive input voltage X is applied, the output is Y = X, while if a negative input voltage X is applied, the output is Y = -X.

- (a) Find $\mathbb{E}Y$.
- (b) Plot the cdf of Y.
- (c) Find the pdf of Y

	Discrete	Continuous
$P\left[X \in B\right] =$	$\sum_{x \in B} p_X(x)$	$\int_{B} f_X(x) dx$
$P\left[X=x\right] =$	$p_X(x) = F(x) - F(x^-)$	0
Interval prob.	$P^{X}((a,b]) = F(b) - F(a)$ $P^{X}([a,b]) = F(b) - F(a^{-})$ $P^{X}([a,b]) = F(b^{-}) - F(a^{-})$ $P^{X}((a,b)) = F(b^{-}) - F(a)$	$P^{X}((a,b]) = P^{X}([a,b])$ = $P^{X}([a,b)) = P^{X}((a,b))$ = $\int_{a}^{b} f_{X}(x)dx = F(b) - F(a)$
$\mathbb{E}X =$	$\sum_{x} x p_X(x)$	$\int_{-\infty}^{+\infty} x f_X(x) dx$
For $Y = g(X)$,	$p_Y(y) = \sum_{x: g(x)=y} p_X(x)$	$f_Y(y) = \frac{d}{dy} P\left[g(X) \le y\right].$ Alternatively, $f_Y(y) = \sum_k \frac{f_X(x_k)}{ g'(x_k) },$ $x_k \text{ are the real-valued roots of the equation } y = g(x).$
For $Y = g(X)$, $P[Y \in B] =$	$\sum_{x:g(x)\in B} p_X(x)$	$\int_{\{x:g(x)\in B\}} f_X(x)dx$
$\mathbb{E}\left[g(X)\right] =$	$\sum_{x} g(x) p_X(x)$	$\int_{-\infty}^{+\infty} g(x) f_X(x) dx$
$\mathbb{E}\left[X^2\right] =$	$\sum_{x} x^2 p_X(x)$	$\int_{-\infty}^{+\infty} x^2 f_X(x) dx$
$\operatorname{Var} X =$	$\sum_{x} (x - \mathbb{E}\overline{X})^2 p_X(x)$	$\int_{-\infty}^{+\infty} (x - \mathbb{E}X)^2 f_X(x) dx$

Table 5: Important Formulas for Discrete and Continuous Random Variables